

Multi-environment trial analysis of count data with complex variance structures using generalised linear mixed models

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GATES *foundation*



- ▶ Motivating multi-environment trial (MET) data

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 - ▶ Factor-analytic Conway-Maxwell Poisson (FA-CMP_μ) model
- ▶ Results & key findings

Motivating MET Data

The motivating MET data is from a series of 13 common bean *Phaseolus vulgaris* variety trials conducted at 9 locations across Ethiopia in the 2022 and 2023 seasons



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Response variable is pod count per plant

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- ▶ Each genotype was replicated 3 times using a row-column design using the `odw` R-package

Genotype concurrence across trials

Trial	AN22	AN23	BK23	HU22	HW23	JM22	JM23	MK22	MK23	PW22	SK22	SK23	WK23	# of columns	# of rows
AN22	160													24	20
AN23	45	135												15	27
BK23	45	135	135											15	27
HU23	16	48	48	48										12	12
HW23	45	135	135	48	135									15	27
JM22	85	35	35	6	35	85								15	17
JM23	35	110	110	23	110	33	110							15	22
MK22	160	45	45	16	45	85	35	160						15	32
MK23	45	135	135	48	135	35	110	45	135					15	27
PW22	85	33	33	4	33	81	33	85	33	85				15	17
SK22	95	24	24	14	24	20	14	95	24	20	95			15	19
SK23	19	48	48	36	48	9	23	19	48	7	17	48		12	12
WK23	37	110	110	23	110	34	101	37	110	33	16	23	110	15	22

Mixed Models

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When the LMM assumptions are violated, an extension to a Generalised linear mixed model (GLMM) is required

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 - ▶ Generalised Poisson distribution
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- ▶ **None of these proposed distributions can account for arbitrarily over and underdispersed count data**

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 1. CMP _{μ} distribution is part of the exponential family
 2. The CMP _{μ} distribution can handle arbitrarily **over and underdispersed data** (Huang 2023)
 3. The dispersion parameter $\nu(\phi)$ is functionally independent of the location parameter μ (Huang & Rathouz 2017)

$$p(y; \mu, \nu) \propto \frac{\lambda(\mu, \nu)^y}{(y!)^\nu}, \quad y = 0, 1, 2, \dots,$$

The mode parameter $\lambda(\mu, \nu)$ is obtained by solving

$$\sum_{y=0}^{\infty} (y - \mu) \frac{\lambda^y}{(y!)^\nu} = 0,$$

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$$\phi \propto 1/\nu$$

Statistical Model

$$l(\boldsymbol{\beta}, \boldsymbol{\gamma}, \boldsymbol{\phi}, \mathbf{y} | \mathbf{u}) = \sum_{j=1}^t \sum_{i=1}^{m_j} \sum_{l=1}^{d_j} \log(p(y_{ijl}; \mu_{ijl}, \nu_j))$$

$$\log(\boldsymbol{\mu}) = \mathbf{X}\boldsymbol{\beta} + \mathbf{Z}_g \mathbf{u}_g + \mathbf{Z}_o \mathbf{u}_o,$$

$$\log(\boldsymbol{\nu}) = \mathbf{X}\boldsymbol{\zeta},$$

- $p(y_{ijl}; \mu_{ijl}, \nu_j)$ is the probability mass function for the CMP_{μ} distribution for the i^{th} genotype in the j^{th} environment (i.e. trial) within the l^{th} replicate block

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- ▶ The vectors $\boldsymbol{\beta}$ and $\boldsymbol{\zeta}$ are each of length t and denote the environment fixed effects for the mean and dispersion parameters respectively, each with corresponding design matrix $\boldsymbol{X}$
- ▶ t is the total number of environments

A factor analytic structure for the random $G \times E$ interaction effect

$$\mathbf{u}_g = (\mathbf{\Lambda} \otimes \mathbf{I}_m) \mathbf{f} + \boldsymbol{\delta}$$

- ▶ $\mathbf{u}_g$ is the vector of random genotype by environment ($G \times E$) interaction effects of length mt , and has a factor analytic structure of order k
- ▶ Denoted as an $FA(k)$ model
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- ▶ $\boldsymbol{\delta} \sim N(\mathbf{0}, \boldsymbol{\Psi} \otimes \mathbf{I}_m)$ is a vector of genetic regression residuals of length mt

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► $\text{var}(\mathbf{u}_g) = (\mathbf{\Lambda} \mathbf{\Lambda}^T + \boldsymbol{\Psi}) \otimes \mathbf{I}_m$

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- ▶ $\mathbf{\Lambda} \mathbf{\Lambda}^T$ captures the 'repeatable' $G \times E$ interaction
- ▶ $\boldsymbol{\Psi}$ is a $t \times t$ diagonal matrix containing the specific variances for each environment

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- ▶ Computationally infeasible for complex variance structures
- ▶ This has been resolved recently with the `glmmTMB` R-package (Brooks *et al.* 2017)
- ▶ `glmmTMB` uses **automatic differentiation** to speed up the computation of high dimensional gradient functions (Griewank and Walther, 2008)

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- The REML adjustment is obtained by marginalising with respect to β as well as $\mathbf{u}$ (Maestrini *et al.* 2024)

$$l_{\text{REML}}(\gamma, \phi; \mathbf{y}) \approx l(\hat{\beta}, \gamma, \phi, \tilde{\mathbf{u}}; \mathbf{y}) + (n - t) \log \sqrt{2\pi} - \frac{1}{2} |\mathbf{H}^*|$$

- $\mathbf{H}^*$ is the Hessian of $l(\hat{\beta}, \gamma, \phi, \tilde{\mathbf{u}}; \mathbf{y})$

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- ▶ Model G×E interaction effects using FA variance structure
- ▶ Model heterogeneous dispersion using the CMP _{μ} distribution
- ▶ Enables partitioning of genetic sources of variation from non-genetic sources and dispersion
- ▶ Goal is genotypic selection by maximising prediction accuracy of $\mathbf{u}_g$

FA models of various order were fit to model the $G \times E$ interaction effects

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- ▶ Random row and column terms fitted for each trial to adjust for spatial field trend through $\boldsymbol{u}_o$

Results & key findings

FA model of order 3 was the best fit as per the AIC criteria

- ▶ FA(3) model explains 75.1% of the total $G \times E$ interaction effects.

$$AIC = -2l_{LA, REML}(\gamma, \phi; y) + 2q$$

G×E variance structure	Number of non-boundary parameters (q)	AIC
Diagonal	49	25857
FA($k = 1$)	62	25695
FA($k = 2$)	74	25694
FA($k = 3$)	79	25688
FA($k = 4$)	89	25695
FA($k = 5$)	94	25701
Unstructured	127	25761

Results of MET analysis

Trial	Trial log Mean $\hat{\beta}_j$	Trial Dispersion $1/\hat{\nu}_j$	Sqrt Genetic Variance $\hat{\sigma}_{g_j}$
AN22	3.29	1.88	0.18
AN23	2.98	1.62	0.13
BK23	2.32	0.72	0.19
HU23	3.00	0.18	0.08
HW23	2.74	0.72	0.10
JM22	2.93	0.69	0.10
JM23	2.63	1.32	0.17
MK22	3.01	1.43	0.17
MK23	3.09	2.51	0.22
PW22	2.19	0.39	0.18
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► All trials have genetic variance

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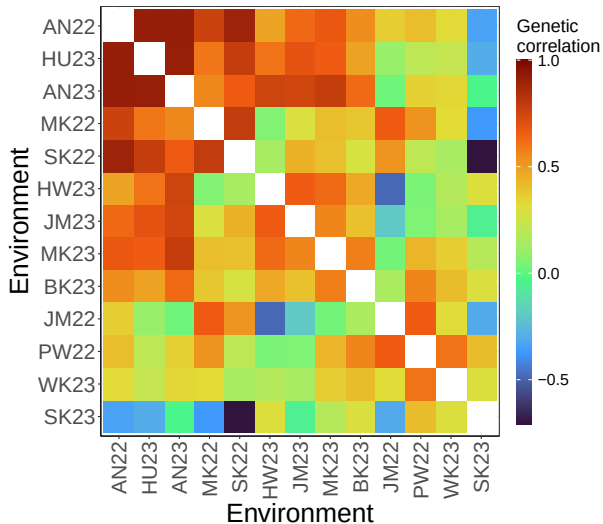
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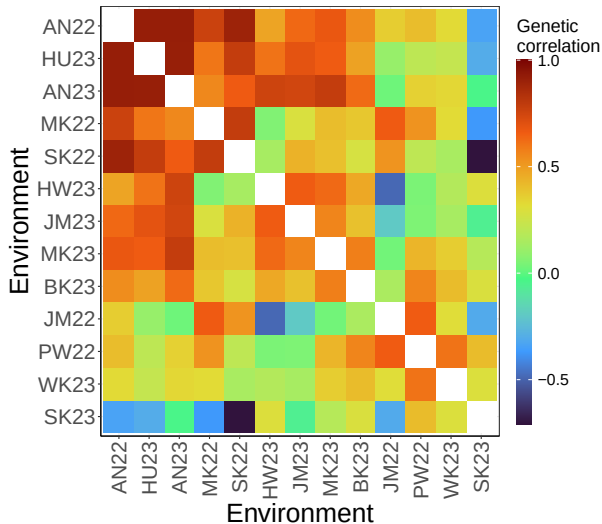
- ▶ All trials have genetic variance
- ▶ Some trials have strong under dispersion
- ▶ Trial with highest mean count is underdispersed
- ▶ Trial with highest dispersion also has the highest genetic variance
- ▶ Trials with next two highest genetic variances are underdispersed

Results of MET analysis - Heatmap



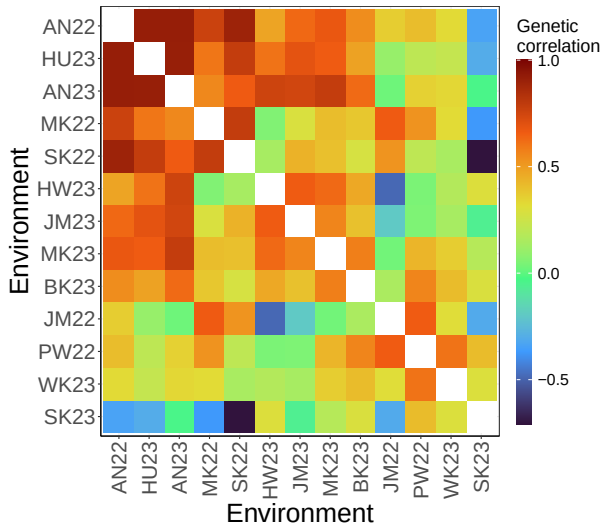
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- ▶ Trials ordered by clustering algorithm

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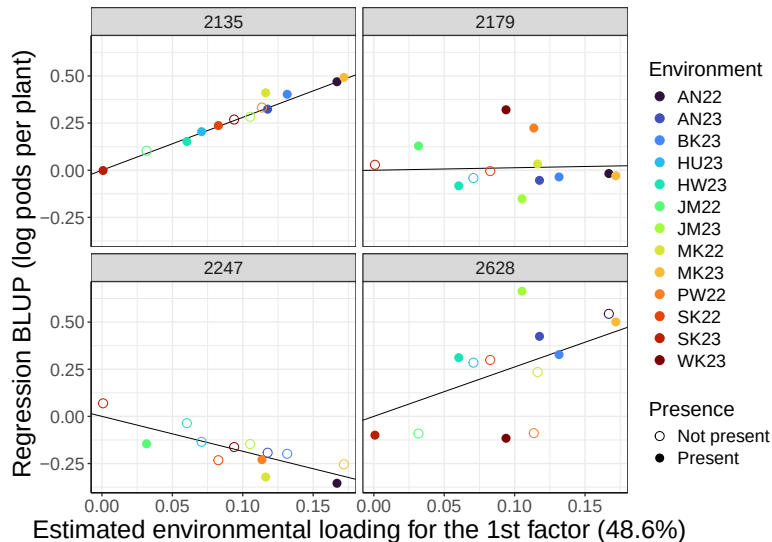
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- ▶ Cluster of trials with high genetic correlation on the top left

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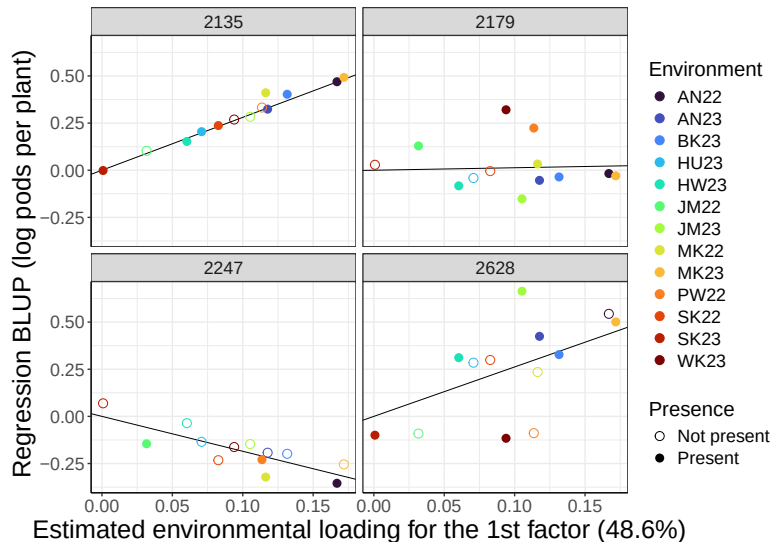
- ▶ Genetic correlations obtained from covariance matrix G_e
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- ▶ Cluster of trials with high genetic correlation on the top left
- ▶ Trials on the bottom right generally have low genetic correlations

Latent regression plot



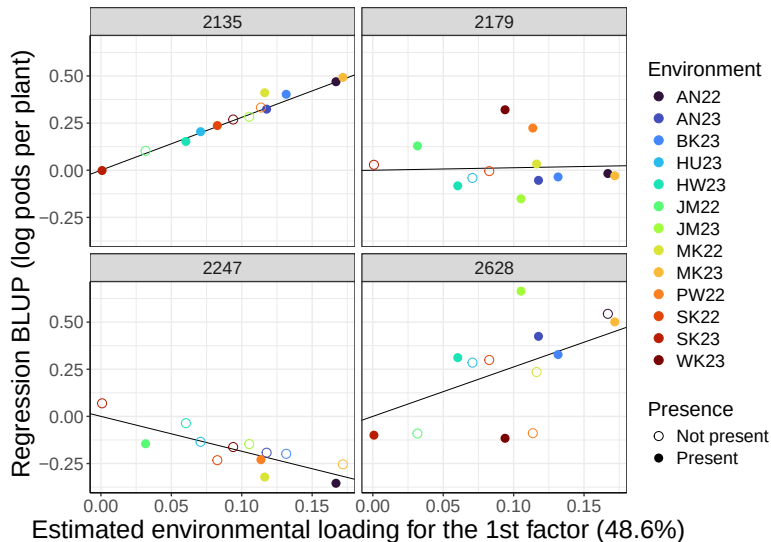
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Latent regression plot



- ▶ Subset of four genotypes
- ▶ All rotated loadings > 0

Latent regression plot



- ▶ Subset of four genotypes
- ▶ All rotated loadings > 0
- ▶ Some genotypes have higher 'lack of fit'

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 - ▶ Automatic differentiation to make model fitting computationally feasible
 - ▶ Laplace approximation estimation method to reduce bias
 - ▶ REML correction to further reduce estimation bias

- ▶ Dr. Alan Huang and Dr. Alison Kelly for their ongoing support, expertise, & mentorship
- ▶ Queensland Department of Primary Industries for financial support & time to work on the project



- ▶ The National Lowland Pulse Research Program at the Ethiopian Institute for Agricultural Research (EIAR) for conducting the trials & offering essential facilities, data & support for this study.
- ▶ The MERCI project team at both EIAR and the University of Queensland within QAAFI, funded by the Gates Foundation



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AUSTRALIA

QAAFI
Queensland Alliance for
Agriculture and Food Innovation

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